Department HW
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HW 25/38

PAPER ON STATISTICS OF REPETITIONS

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## Statistics of Repetitions

In order to be able to obtain reliable estimates of the value of given repeats we need to have information about repetition in plain language. Suppose for example that we haveplaced two messages together and that we find repetitions consisting of two a tetragramme, two bigrammes and fifteen singel letters, and that the total 'overlapt was 105, i.e. that the maximum possible number of repetitions which could be obtained by altering letters of the messages is 105; suppose also that the lengths of the messages are 200and 250; in such a case what is the probability of the fit being right, no other information about the days traffice being taken into consider tion, but information about the character of the unenciphered text being available in considerable quantity?

In theory this can be solved as follows. e take a wast number typical dedodes of mesmages, sav 10'0 , and from these we select alllof length 200 and all of length 250. We encipher all of these messagew at all possible b sitions on the machine (neglecting for simplicity the complication due to different daily keys). We then compare each message 200 long with each 250 long in such a way as to get an overlap of 105 as with the fit under consideration. From the resulting comparisons we pick out just those cases where the repetitions have precisely the some form as the case in question. This set of comparisons will be called the 'relevant' comparisons. Among the relevent comperisons there will be some which are 'right' comparisons, i.e. where corresponding letters of the two messages were enciphered with the same position of the machine. The probability that our original fit was right can now be expressed in the form

Number of right relevant comparisons
Total number of relevant comparisons

The work involved in this theoretical method can be vistly reduced if we make a few harmless as umptions. In the first place if we assume that the encipherment keys at the various positions of the machine are hatted we can calculate the number of relevant wrong comparisons. Suppose the total number of repeated letters in the ase in question is R, then

Number of relevent wrong comperisons = (1) R (26) R (26)

For the calculation of the number of relevant right comparisons we have to make other assumptions. The sort of assumption that we heed is that a repetition in one place is not made any the more or less likely by a repetition elsewhere. Actually this assumption would not be quite true, as it clearly does not hold in the case of a jacent letters. For most practical purposes I think the following essumption is sufficiently near to the truth }-If we know that at a certain point P there is not a repetition, th en knowledge that there is or is not a repetition at a point A before P does not make a repetition at a point B after P either more likely or less likely. With this as a motion xxxx xunldxgetxthexrightxdistribationxufxresotitionxerrangementax MENNEYARE The probability of a repetition at any point is also independent of its distance from the ends of either message. With these assumptions we could get the right distribution of numbers of comparisons between ther various repetition figures if we assume the repetition figures for the comparisons constructed in this way. We are given an urn containing a number of xitox efrozper cards, some bearing the words 'no repeat', some 'simple repeat', some 'bigramme', some'trigramme', and so on. To construct a rendom semple of repetition figures for comparisons of given length we make a series of draws from the

The first few draws determine the recetition figure for the comperison first maxing, the next few for the next mixing comperison, and so on. When we draw 'no repeat' we have to add a 'o' to the repetition figure, when we draw 'simple repeat'we add '%o', for 'bigramme' we add 'xxo' and so on. When we have got to the right length of overlap required the comparison is completed and our next draws refer to the next comparison. Then If it happens that the right length is never reached because we'jump past it' then we scrap that comparison, and go on to the next. As an example suppose that we are making comparisons with an overlap of 12, and that our first draws are 'tetragramme', 'no rep', 'no rep', 'bigramme', 'no rep', 'trigramme', 13- gramme, then 'no rep' 13 times, our first two comparisons will have the repetition figures

XXXX0000XXX00

00000000000

Two problems arise from thes picture

- 1) How do we calculate the correct proportions of cards in the  $\operatorname{urn}\nolimits ?$
- 2) Given the proportions of the cards in the urn, how do we calculate the number of right relevant comparisons, and hence the probability of a given fit?

The correct propertion of the cords in the urn can be calculated from the actual distribution of repetitions in the case of messages correctly set, or, what comes to the same thing, in messages unenci hered and arbitrarily set. Let us suppose that we have a large number of such comparisons of unenciphered messages, and that the messages are sufficiently long that complications arising from the ends of the messages can be neglected. The propertion of cards bearing the words 'simple repeat', 'bigramme', 'trigramme' etc must obviously be in the same ratio as the number of corresponding repeats in our comparisons. The number of 'no repeat' cards will be calculated slightly differently as we have to subtract one case of 'no repeat' for each sequence of repeating letters.

To get the best value from given material we naturally make every postible comperison. If we do this the right number of repetitions can b e calculated quite easily without actually making the comparisons. Theoretically we can imagine the complete set of comperisons made in this way. First of all we write out all the decodes (sey 50 of them) one after another round a circle: suppose that the number of letters on this circle is N. The whole isthen repeated on a concentric circle. All pos ible comparisons can be made by rotating the one circle with respect to the other. From the ene the ese we have to remove the comparison in which the circles are not rotated at all, for obvious reasons. Also when the rotation is more the n 180° we get essentially the same comparison as one with less than 180°. The net effect of this, taking into account also the EDE special case of exact 180° rotation, is that the total overlap of all the comparisons of  $\frac{12}{2} \frac{N(N-1)}{2}$ . Now let us consider for

example the total number of tetragram e repeats in all these comparisons. These can be divided into the repeats arising from AAAA those from AAAB .... those from ZZZZ, the lergest contribution arising presumably from such tetrogram as as EINS. The number of tetragrammes arising from EINS consists of the n umber of pairs of hexagrammes such as QEINSR, VEINSW in which th e first letters of each are different, the last different. and the remainder spell EINS. This number of pairs we will call the 'actual number of tetragram e repeatsarising from EINS8. The 'adtual number of tetragramme repeats' is obtained by summing over AAAA, AAAB, ..., EINS, ... ZZZZ. This 'actual' number is not easily calculated directly, but we can more easily obtain the 'apparent number of tetragramme repeats', and this leads to the actual number. The 'apparent number of tetragramme repeats arising from EINS' is defined to be the number of pairs of occurences of EINS in the material, and Tthe apparent number of tetragramme repeats' defined by summation. We can also define the apparant number of testragramme repeats in a comparison as the numbers of different series xxxx in the comperison. Thus a heptagramme repeat gives four apparent tetragramme repeats. The actual number of repeats can be calculated from the apparant in this way. Let Mr be the apparent number of r-grammes, and m Nr th e actual number. Then

So that

Mr. Mr.: Nr. + Nr. + Nr. + Nr. + ...

Nr: (Mr. Mr.) - (Mr. - Mr.) : Mr. - 2 Mr. + Mr.

It is therefore sufficient to calculate only apparent numbers and to carry these two stages further that we went to go with the actual numbers. In practice octagramme repeats are so certain to be right that it will be sufficient toheve statistics only as fer as heptegrammes. We therefore need statistics of mater apparant numbers of reperts as for as \$-grammes. To get these numbers of apparent repeatsit is sufficient to take all the 9-grammes in the material (i.e. on the circle) and to put them into alphabetical order. This can be done very cojveniently by Hollerith. The number of trigremme repeats can then be found very simply ( lthough with a good de 1 of labour) by bek considering only the first three letters of each 9-gramme. Suppose we renote hax by t a typical trigram erand by M, the n umber of its occurrences, then the apparent number of trigramme  $\sum_{t} \frac{n_{t}(n_{t}-1)}{2}$ repestsis

In our leter conculctions it is convenient elso to regard the comparisons in the wrong places as elso constructed by drawing from moth or urn. The proportions in this urn can be in theory be calculated in the same way but from massages consisting of hatted series of letters. In this case the proportion apparent manner of togrammes is the so that the obtual proportion of togrammes is  $\frac{2^{5/2}}{2^{1/2}}$ .

When calculating the proportions of cards in the urn we must remember that the total number of cards is not  $\frac{N(N-1)}{2}$  but is less than this by  $\sum v N_{ic}$ .

In our later calculations it is convenient to regard the comparisons in wrong places as also constructed by drawing from an urn. In this case we easily set that the apparent number of r-grammes is  $\frac{N(N-1)}{26}$ , and from this we deduce that the actual proportion of r-gramme cords is  $\frac{25}{26}$ , and of no repert cards is  $\frac{25}{26}$ .

We now turn to the proble | of c lulating th e probability of a given fit when we know the proportion of ragramme cards in the urn for each r. The calculation is going to be slightly complicated by the convention which we introduced, that not all drawings can lead to a comparison. We have therefore to c loulate the proportion of draws which lead to a comparison, i.e. in which the length does not overshoot the mark. Th e endwar is that as he length of overlab tends to infinity the proportion t nds to ; in the case of hatted material this is 25

Now put  $A = 1 - Z_{\mathcal{L}_{\nu}}$  . Consider repetition figure in which The proportion of right draws which are felevent is

and the proportion of the right comparisons which are relevant is (essuming L resonably 1 rge)

Similarly with the urn whose proportions were made up from hatted material we find for the comperisons proportion of wrong draws which are relevant

Hence the odds\*on our fit are

when odds on our fit are
$$\gamma : \lambda \frac{26 - (14 \overline{\Sigma} r_{Kr})}{26} \left( \frac{26 \beta}{26} \right)^{L+1} - \overline{\Sigma} (r+1) k_r \frac{\infty}{\Gamma_{L,1}} \left( \frac{26^{r+1} \alpha_r}{26} \right)^{k_r}$$

where  $\lambda$  is the priori odds. This is most conveniently written as

$$\log q = \log \lambda + \sum \mu_r k_r - \nu L + \log (1 - \sum k_r)(1 + \sum r k_r)$$
where  $\mu_r \cdot \log \frac{k_r^2 b^{2d}}{25} - (r^2) \log \frac{26A}{25} = nd$ 
 $\nu \cdot \log \frac{26}{26A} = \sum k_r - \frac{2}{5}$ 

<sup>\*</sup>The odds on an event are defined to be the are abili y of the event divided by the probability of its negation

In the case of overlap zero there is a discrepancy of  $\frac{L_2(1-\widehat{Z}_{K_r})(1-\widehat{Z}_{K_r})}{(1-\widehat{Z}_{K_r})(1-\widehat{Z}_{K_r})}$  due to the overlap not being long. This term is in any case microscopic.